Response to the First Referee.

First, we would like to thank the Referee for her/his comment.

The referee has made a point that the similarity report is too high even for a review article. We have addressed this issue, rewriting several portions of the manuscript, referring to the relevant references where necessary, and dropping some paragraphs not necessary for this review’s discussion. We hope that the manuscript is now in a more satisfactory state.
Response to the Second Referee.

First, we would like to thank the Referee for her/his comments, all of which we have attempted to address. We think that the paper has been improved by them. Now we detail our response to each comment.

1. Despite the intuitive, valid fractal analysis, I feel that the manuscript does not have many new elements to showcase. The reference to turbulence in efforts to physically connect solar, interplanetary, and magnetospheric timeseries is biased in its framework. The manuscript effectively shows the effect of intermittency in the fractal dimension of timeseries, regardless of turbulence. Intermittency is a term that is broader than turbulence: turbulent timeseries may be intermittent, but not all intermittent timeseries stem from turbulent systems.

The scatter diagram of Figure 1 creates some “dust-like” fractals in case of intermittency (in this case, storm-time dips in Dst). Dust-like structures typically give rise to a fractal dimension smaller than Dmax=1, where Dmax is the embedding (i.e., Euclidean) dimension of the studied space. In Figure 1, Dmax=2, hence the dust-like structures in the lower-left part of the image show a fractal dimension $D < 1$ (see, e.g. Schroeder, M.: Fractal, Chaos, Power Laws. Minutes from an Infinite Paradise, Freeman, New York, NY). If no significant intermittency is present, one is left with the upper right part of Figure 1 that typically gives $1 < D < 2$.

Interpreting intermittency in general as turbulence and drawing physical conclusions from it is the main drawback of the manuscript.

It was not our intention to force a connection between all three systems studied through the concept of turbulence. As stated in the manuscript, the GOY shell model has been shown to exhibit dissipative events whose distribution follows the same power-law statistics as observed in turbulent magnetized plasmas (Boffetta et al., 1999; Lepreti et al., 2004; Carbone et al., 2002), and our main goal was to test whether such
bursty behavior exhibited fractal features similar to those found in the
$\text{Dst}$ analysis.

Although various works suggest the presence of turbulence in the Earth’s
 magnetosphere, the question of the validity of the GOY shell model to
describe such phenomena is far beyond the scope of our paper. As the
referee correctly points out, it is the intermittency of the time series
which is the relevant feature, and in fact that is what justifies the use
of certain values of $\nu$ and $\eta$ for the simulations, since, in general, inter-mittency levels similar to the $\text{Dst}$ timeseries are not observed for
arbitrary values of these parameters.

We have attempted to clarify this issue in various parts of the text. For
instance, in the second paragraph of Sec. 4.

The new text reads:

_We first notice that, in general, setting parameters $\nu$ and $\eta$ with arbi-
trary values yields $e_0(t)$ series which do not have the necessary inter-
mittency level to resemble the $\text{Dst}$ time series. Compare, for instance,
the different panels in Fig. 16 in Domínguez et al. (2017), which shows
that $Pm = 0.2$ leads to a very noisy output, unlike simulations with
$Pm = 1.0$ or 2.0, where individual, large peaks can be easily identi-
fied from the background. In fact, previous studies have shown that the
statistics of bursts follows a power law for $Pm = 1$ (Boffetta et al.,
1999; Lepreti et al., 2004; Carbone et al., 2002), and for this reason
we start by taking $Pm = \nu/\eta = 1$. _

Also, in the final paragraph of the same section.

The new text reads:

_Results suggest that the intermittency level of the output time series is
relevant, which has led us to perform the analyses for the shell model
within a certain range of values of the Prandtl number, as well as of
the viscosity and resistivity._

2. This leads to insufficiently justified conclusions such as the
correlations between D from $\text{Dst}$ timeseries and the solar flare
Indeed, there is connection if an eruptive flare (flare + coronal mass ejection) leads to a magnetospheric storm within 1–3 days. However, the correlation seen in Figure 9 is not due to physics but due to the fact that any two intermittent timeseries with intermittent excursions roughly matching in time will show similar correlations. I am afraid this is a common fallacy, appearing in several interdisciplinary studies of timeseries giving, not surprisingly, incidental correlations.

Thanks for pointing out this issue. The figure that the referee mentions (Fig. 9 in the previous version of the manuscript), was part of an exploration of possible correlations between fractal dimensions and various indices, using solar and geomagnetic timeseries, which was made in Domínguez et al. (2014). As the referee says, a better statistical and physical analysis is needed to state whether these correlations hold or not. Besides, in the context of the present manuscript, it is not a relevant discussion, since we focus on the series themselves, not on their correlations with others. We have thus dropped Fig. 9 in this version of the manuscript.

3. Another unjustified conclusion is the one drawn from Figures 7, 8, namely that “results suggest that the box-counting dimension consistently decreases when the storm approaches” (p.8; top). However, the decrease is not due to the storm but due to the pre-storm disturbances (hours > 1400 and up to the storm’s onset). These disturbances are not necessarily related to the storm. Similar disturbances appear at times < 500 hours in the absence of a storm. Not surprisingly, D in this interval is very similar to the pre-storm D that is indeed decreasing. Again, it is the (most likely incidental, as it starts ~300 hours prior to the storm) minor intermittency in the timeseries that causes the decrease in both cases, regardless of the storm. Finding a unique pre-storm signature is the challenge here and the manuscript does not seem to contribute significantly to this cause.
We agree that conclusions need to be toned down, and that the present analysis cannot suggest that the decrease observed before the storm is related to the storm itself. However, our aim in this manuscript is focused rather on the dissipative events themselves and the fractal dimension, not on the finding of precursors for geomagnetic activity, an issue which requires further, detailed analysis.

Thus, we have changed the wording in the sentence mentioned (now at the bottom of page 5).

The new text reads:

As shown in Domínguez et al. (2014), the box-counting dimension of the Dst index decreases as the storm approaches for all cases studied. Moreover, this decrease occurs before the window includes the geomagnetic storm, as marked by the vertical lines in Fig. 5. Whether this is relevant for forecasting geomagnetic storm needs further study, as it may simply be due to an increase of the intermittency in the time series, unrelated to the upcoming dissipative event.

4. The above issues render the penultimate conclusion of the manuscript (p.14) also biased. I see no point in re-doing the analysis unless more physical and statistical arguments for the apparent correlations are used alongside the analysis of the fractal dimension.

We have indeed performed more systematic analysis than the ones mentioned in this manuscript, but were left in the cited references and not included in the current text.

Cross correlation analyses between the Dst timeseries and its fractal dimension were performed. This is not a direct calculation, as both time series have different resolutions, and thus interpolation of the fractal dimension time series is needed to match the resolution of the geomagnetic index. This analysis was made for individual storms and full year data, and is included in Domínguez et al. (2014). This is mentioned in the final paragraph of Sec. 3.

On the other hand, $p$-value analyses were systematically done for the
shell model simulations, for a wide range of values of $\nu$ and $\eta$, considering $P_m = 1$ and $P_m \neq 1$. This allowed us to find a range of values of the simulation parameters where the correlation between $\epsilon_b(t)$ and its fractal dimension is statistically significant.

We have mentioned this issue in the first paragraph of page 8, relating it to the problem of intermittency.

The new text reads:

*We first notice that, in general, setting parameters $\nu$ and $\eta$ with arbitrary values yields $\epsilon_b(t)$ series which do not have the necessary intermittency level to resemble the Dst time series. Compare, for instance, the different panels in Fig. 16 in Domínguez et al. (2017), which shows that $P_m = 0.2$ leads to a very noisy output, unlike simulations with $P_m = 1.0$ or 2.0, where individual, large peaks can be easily identified from the background. In fact, previous studies have shown that the statistics of bursts follows a power law for $P_m = 1$ (Boffetta et al., 1999; Lepreti et al., 2004; Carbone et al., 2002), and for this reason we start by taking $P_m = \nu/\eta = 1$.\*\*

And again in the last paragraph of Sec. 4.

The new text reads:

*In Domínguez et al. (2017) a more detailed analysis is carried out on the shell model results, exploring other simulation parameters ($\nu$, $\eta$, magnetic Prandtl number), other criteron for defining active states ($n = 5$), and a systematic study of the correlations between the fractal dimension and the occurrence of dissipative events by means of the Student’s $t$-test. Results suggest that the intermittency level of the output time series is relevant, which has led us to perform the analyses for the shell model within a certain range of values of the Prandtl number, as well as of the viscosity and resistivity.*
Response to the Third Referee.

First, we would like to thank the Referee for her/his comments, all of which we have attempted to address. We think that the paper has been improved by them. Now we detail our response to each comment.

1. The similarity index of 28% could be acceptable for a review provided that all credits are given, even if the authors of the previous published papers are also on the authors list of the review. But 26% (including Figures 1 and 2) are simply copied from Dominguez et al. (2014).

We have made several modifications in various parts of the manuscript in order to deal with this issue, including dropping parts of the text that were not relevant for the line of the discussion intended in this paper.

2. Obviously, as mentioned in the introduction fractal dimensions have already often been calculated for space and laboratory magnetized plasmas in nature, including the magnetosphere (e.g., J. Geophys. Res. 96, 16031, 1991) and the solar wind (e.g., J. Geophys. Res. 114, A03108, 2009; Astrophys. J. Lett., 793:L30, 2014). But the subject of the submitted review is rather limited to very selected examples of space plasmas, basically only to geomagnetic activity (besides preliminary results applied to magnetic clouds and additional discussion in the context of the turbulence shell model) and therefore the title of the review should possibly be much more specific.

We have changed the title to “Evolution of fractality in space plasmas of interest to geomagnetic activity”, in order to be more specific and consistent with the content of the manuscript.

3. By the way, the phenomenological MHD shell model describes the energy cascade in turbulence that sometimes exhibits fractal characteristics, but geomagnetic storms have quite different more intermittent characters, sometimes related to multifractality. It would be nice to provide convincing physical
arguments justifying application of this model to dynamics of geomagnetic activity.

Maybe we should stress that we are not attempting to use the MHD shell model to account for $D_{st}$ dynamics. Our interest in the connection between two model arises from the possibility of having similar intermittent behaviors, as the shell model can also yield simulations which do not exhibit intermittency levels which resemble the $D_{st}$ time series.

We have added a text in the first paragraph of page 8, related to this issue.

The new text reads:

We first notice that, in general, setting parameters $\nu$ and $\eta$ with arbitrary values yields $\epsilon_0(t)$ series which do not have the necessary intermittency level to resemble the $D_{st}$ time series. Compare, for instance, the different panels in Fig. 16 in Domínguez et al. (2017), which shows that $P_m = 0.2$ leads to a very noisy output, unlike simulations with $P_m = 1.0$ or $2.0$, where individual, large peaks can be easily identified from the background. In fact, previous studies have shown that the statistics of bursts follows a power law for $P_m = 1$ (Boffetta et al., 1999; Lepreti et al., 2004; Carbone et al., 2002), and for this reason we start by taking $P_m = \nu/\eta = 1$.

We have also been careful in the use of words, refering to dissipative events in the shell model as “active” states, whereas in the $D_{st}$ time series they correspond to “storm” states, with definite physical meaning.

The possible connection between geomagnetic activity and the GOY shell model has been suggested in Lepreti et al. (2004), but testing this goes beyond the simple fractal analysis we propose in this manuscript.

4. page 3, lines 16-18: Admittedly, there is no commonly accepted definition of a fractal (for example, according to B. B. Mandelbrot, 1977: “a fractal is by definition as set for which the Hausdorff Besicovitch dimension strictly exceeds the topo-
logical dimension”). But certainly, “noninteger numbers measuring the complexity” is rather unclear (maybe roughness, irregularity) and certainly not general (e.g., for the trail fractal Brownian motion its fractal dimension is integer, equal to 2, but greater than 1, the topological dimension).

We agree with the Referee in that one has to be careful with definitions. However, we should notice that the cited sentence in our paper refers to the problem of defining fractal dimensions, rather than fractal objects. So, for a given fractal object, there are several ways to define its dimension, and this is what we intended to stress. We have modified the sentence to be more clear.

The new text reads:

_In general it can be said that they are numbers, which can be non-integer, measuring the complexity of a data set._

5. Section 2: The methods of nonlinear time series are well-known, see e.g. the textbook of H. Kantz and T. Schreiber published by Cambridge University Press in 1997. Besides the box-counting (zero-order, capacity) dimension one can also define the (higher-order) generalized dimensions (related to a multifractal spectrum), which are (e.g., the correlation dimension) much more suitable for nonlinear dynamical systems as is in the case of the magnetosphere. Therefore, I would like to ask why the authors use only the box-counting method, which is certainly not very reliable?

Our aim was specifically to investigate whether a single fractal dimension may yield useful information on the systems studied, and in what sense. Certainly, given the complexity of the system, there is no guarantee that this is possible at all, but we have found some positive results as described in the manuscript, which we think are interesting. Other choices for that single fractal dimension could have been made. However, rather than changing the type of dimension used, we think it is
more interesting to perform a multifractal analysis, in accordance with the nature of the systems studied, and this is currently in process.

This is mentioned in the final paragraph of Sec. 7.

The new text reads:

*Given the rich and complex dynamics governing the evolution of magnetized plasmas, we would not expect that a single index would be able to capture all their relevant information. In fact, multifractal analysis should be made in order to represent the dynamics of the systems studied more accurately, and such an analysis is currently being prepared for future publication.*

6. Further, for estimation of any fractal dimension one would require at least approximate stationarity. Hence, my main question is how do the authors cope with non-stationarity of the data under their study, especially during storms. I think that in the magnetospheric studies it would be more difficult task than in the case of the solar wind plasma. Maybe also some filtering is needed before estimating the actual dimension of the fractal structure (see, e.g.: Phys. Rev. E 47, 2401, 1993; Physica D 122, 254, 1998).

It is not clear that, for the kind of analysis we are interested, stationarity is a requisite to get meaningful results. For instance, the magnetic cloud analysis clearly involves a process where various degrees of stationarity are found. Without looking at the fractal dimension, one could argue that the flux rope stage satisfies the stationarity criterion, the sheath does not, and the solar wind stages could also be approximately stationary. And yet, calculation of the fractal dimension on each state, regardless of its level of stationarity, yields useful results, being able to distinguish the various stages.

This is because the fractal dimension that we calculate is related to the intermittency level of the time series, which is also why storms leave a signature in the dimension, a signature which could be lost with filtering, as suggested by Fig. 8 in our paper. The issue of the need
for stationarity in the *Dst* or shell model time series should be studied more systematically in order to give a definitive answer.

7. Results and Conclusions: Relation of the fractal dimensions to storms should be better justified. Namely, a decrease of the fractal dimension based on *Dst* index presented in Figures 8 and 9 during storms may simply artificially result from lack of stationarity. Anyway, a more comprehensive nonlinear time series analysis is needed before drawing any robust conclusion (e.g., page 13, line 8ff).

We have attempted to tone down the conclusion in this respect. Figure 9 of the previous manuscript has been dropped from the current version of the manuscript, since it was not relevant to the main discussion. Regarding Fig. 8 in the previous version (Fig. 5 in the current one), it is true that the decrease in the fractal dimension previous to the storm could be due to pre-storm intermittency unrelated to the upcoming storm. However, our aim in this manuscript is focused rather on the dissipative events themselves and the fractal dimension, not on the finding of precursors for geomagnetic activity, an issue which requires further, detailed analysis.

Thus, we have changed the wording in the sentence mentioned (now at the bottom of page 5).

The new text reads:

*As shown in Domínguez et al. (2014), the box-counting dimension of the *Dst* index decreases as the storm approaches for all cases studied. Moreover, this decrease occurs before the window includes the geomagnetic storm, as marked by the vertical lines in Fig. 5. Whether this is relevant for forecasting geomagnetic storm needs further study, as it may simply be due to an increase of the intermittency in the time series, unrelated to the upcoming dissipative event.*

Also, the “Some robust behaviors are identified” sentence in the conclusions has been dropped, in order to moderate the conclusions.
Evolution of fractality in magnetized-space plasmas of interest to geomagnetic activity

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Abstract. We studied the temporal evolution of fractality for geomagnetic activity, by calculating fractal dimensions from Dst data and from an MHD shell model for a turbulent magnetized plasma, which may be a useful model to study geomagnetic activity under solar wind forcing. We show that the shell model is able to reproduce the relationship between the fractal dimension and the occurrence of dissipative events, but only in a certain region of viscosity and resistivity values. We also present preliminary results of the application of these ideas to the study of the magnetic field time series in the solar wind during magnetic clouds. Results suggest that, which suggest that it is possible, by means of the fractal dimension is able, to characterize the complexity of the magnetic cloud structure.

Copyright statement. TEXT

1 Introduction

The nontrivial interaction between the Sun’s and the Earth’s magnetosphere There is a nontrivial magnetic interaction between Sun and Earth, coupled by the solar wind, leads to a rich variety of phenomena which has attracted interest to the study of space plasmas for decades, and more recently to the possibility of forecasting of space weather, an issue of large relevance in our increasing technology-dependent society.

Various models and techniques have been developed to study the plasma behavior in the Sun-Earth system. Of these, the study of complexity has been of great interest, as they are it is capable of providing new insights and reveal possible universalities on issues as diverse as regarding universal behavior related to geomagnetic activity, turbulence in laboratory plasmas, physics of or the solar wind, among others to name a few (Dendy et al., 2007; Klimas et al., 2000; Takalo et al., 1999; Chang and Wu, 2008; Valdivia et al., 1988). In particular, these studies have shown that systems such as the it has been suggested that various magnetized plasma systems are in a self-organized critical state, exhibiting fractal and multifractal
features which relate them to a broader class of complex systems. This has been the case in studies on the Earth’s magnetosphere (Chang, 1999; Valdivia et al., 2005, 2003, 2006, 2013), the solar wind (Macek, 2010), the solar photosphere, and solar corona (Berger and Asgari-Targhi, 2009; Dimitropoulou et al., 2009), are in a self-organized critical state, and exhibit complex features such as fractality and multifractality. Some authors have discussed the relationship between the fractal dimension as a measure of complexity, and physical processes in magnetized plasmas in the Sun-Earth system, including the possibility of forecasting geomagnetic activity (Aschwanden and Aschwanden, 2008; Uritsky et al., 2006; Georgoulis, 2012; McAteer et al., 2005, 2010; Dimitropoulou et al., 2009; Conlon et al., 2008; Chapman et al., 2008; Kiyani et al., 2007).

In our work we use the box-counting fractal dimension (Addison, 1997), because of its simplicity and which is as simple a measure of complexity, and with an intuitive meaning. Certainly, a single fractal dimension cannot provide all information on complexity for arbitrary systems, in particular if they also exhibit multifractal behavior as well, as expected in systems in general. Moreover, most systems of interest also have multifractal features, such as the magnetospheric system (Chang, 1999), models of turbulence (Kadanoff et al., 1995; Pisarenko et al., 1993), and the solar wind (Chapman et al., 2008); but it is interesting to note that it does describe some relevant features of these time series’ complexity, as it has been successfully used in previous works relevant to the Sun-Earth system (Osella et al., 1997; Kozelov, 2003; Gallagher et al., 1998; Georgoulis, 2012; Lawrence et al., 1993; Cadavid et al., 1994; McAteer et al., 2005). Furthermore, we will thus use the box-counting dimension as a fast approach to systematically study our systems of interest, and a first step to detect universal features worth of further study.

It is also worth noting that the fractal dimension we calculate is. Besides, we will calculate fractal dimensions based on a scatter diagram (see e.g. Witte and Witte, 2009), whereas previous studies have been done with other Witte and Witte (2009), unlike some previous studies, where different methods or data (Kozelov, 2003; Uritsky et al., 2006; Balasis et al., 2006; Dias and Papa, 2010) were used.

These ideas were implemented by us in Ref. (Domínguez et al., 2014) Domínguez et al. (2014) to study the Dst time series and solar magnetograms, and the possible correlation between solar and geomagnetic activities as evidenced by the box-counting fractal dimension. Individual events, complete years of high geomagnetic activity, and the full 23rd solar cycle were studied with this technique, successfully finding that the fractal dimension, and more specifically its evolution, has —despite its simplicity— relevant information on the complex behavior of these systems and their eventual correlation.

Results above were robust, in the sense that they were observed across a wide range of time scales, which suggests that any model describing the dynamics of geomagnetic activity should reproduce a similar fractal behavior. This is our motivation to study a shell model for MHD turbulence within this framework.

Evidence of turbulence in the Earth’s magnetosphere has been found by various spacecraft observations (Nykýri et al., 2006; Sundkvist et al., 2005; Zimbaro et al., 2008), and several authors have studied magnetospheric MHD turbulence (see, e.g., Borovsky (2004); Hwang et al. (2011); El-Alaoui et al. (2012)). One interesting approach has been the proposal of analytical models depending on few. However, given the large number of degrees of freedom, which nevertheless retain relevant simulation of turbulent systems has a large computational cost, which has led to the development of analytical models which, while sharing
statistical properties of the magnetospheric behavior, such as the power-law distribution and multifractal features of dissipative events. \textit{Systems under study,} depend only on a few degrees of freedom (Chapman et al., 1998; Valdivia et al., 2006).

Shell models constitute an intermediate level between such these models and first principles approaches. They are low-dimensional models, based on a system, we find shell models, consisting of a set of coupled equations mimicking which are similar to the spectral Navier-Stokes equation, and have been but which are also low-dimensional models. They have been successfully used to describe turbulence in magnetized fluids, being able to deal with large Reynolds numbers without the associated computational cost of simulations based on first principles, nonlinear fluid equations (Ditlevsen, 2011), and describing the main statistical properties of magnetohydrodynamic (MHD) turbulence (Chapman et al., 2008), without the computational cost of performing high Reynolds numbers simulations directly from the fully nonlinear fluid equations (Ditlevsen, 2011).

Dissipative--In fact, it has been shown that dissipative events in shell models have been shown to follow can be taken to represent solar flares, and that their distribution follows the same power-law statistics of observed events as observed in turbulent magnetized plasmas, as found in Refs. Boffetta et al. (1999); Lepreti et al. (2001); Carbone et al. (2002), where dissipative events in the model were taken to represent solar flares. In fact, these works suggest that (Boffetta et al., 1999; Lepreti et al., 2000).

As suggested in Lepreti et al. (2004), flares and geomagnetic activity should be the result could be the results of dissipation bursts within a turbulent environment (Lepreti et al., 2004).

In a previous work (Domínguez et al., 2017), we have applied the box-counting fractal dimension to study the complexity in an MHD shell model, analyzing the correlation between it and the energy dissipation rate, showing that, for certain values of the viscosity and the magnetic diffusivity, the fractal dimension exhibits correlation with the occurrence of bursts, similar to what had been found with geomagnetic data (Domínguez et al., 2014). This suggests that shell models do not only reproduce the power-law statistics of dissipative events in turbulent plasmas, but also some features of its fractal behavior.

In this manuscript we review our results in this field, where the fractal dimension is calculated in order to measure complexity in magnetic field time series. The method is used to characterize the occurrence of time series is measured by means of the fractal dimension. Thus, we characterize events such as geomagnetic storms by means of analyzing the Dst time series in various time scales (described in Secs. 2–5, and discussed previously in more details in Domínguez et al. (2014)), and the occurrence of dissipative events in an MHD shell model simulation (Sec 4–5), see Domínguez et al. (2017) for more details). We also present preliminary results dealing with spacecraft data for the solar wind, related to the appearance of magnetic clouds (Muñoz et al., 2016) (Sec. 5).

2 Fractal dimension

We are interested to estimate the fractal dimension to various time series for magnetic data. We now explain the method, using as an example the hourly Dst time series (World Data Center for Geomagnetism, Kyoto).

There are various ways to define a fractal dimension. Fractal dimensions can be defined in various ways in general, and for a time series in particular as well (Addison, 1997; Theiler, 1990). Although there is no simple way to relate different definitions, in general it can be said that they are noninteger numbers, which can be non-integer, measuring the complexity of

3
a data set. In this work, we estimate the fractal dimension using the box-counting method (Addison, 1997) as we now describe as shown below. First, we construct a scatter diagram for each $Dst$ time series. If $Dst^i$ is the $i$-th $Dst$ datum in the series and $N$ is the total number of data, the scatter diagram is a plot of $Dst^{i+1}$ versus $Dst^i$, for $1 \leq i \leq N - 1$, as shown in, by plotting each datum versus the next one (see Fig. 1).

![Figure 1](image)

**Figure 1.** Scatter diagram for the hourly $Dst$ time series corresponding to the first storm state ($using data from 6 to 20 March$, 1989. (More details in see 1989, containing a large geomagnetic storm. (Taken from Domínguez et al. (2014).) The size of the square box is $\epsilon$.

Then, the scatter diagram is divided into square cells of a certain size $\epsilon$, and we count the number $N(\epsilon)$ of cells which contain a point belonging to the set. Finding the range of values of $\epsilon$ by decreasing $\epsilon$, we eventually find a region where the number of cells containing points scale as a power law with $\epsilon$ where $\log(N(\epsilon))$ scales linearly with $\log(\epsilon)$, the scatter diagram:

$$N(\epsilon) \propto \epsilon^{-D}$$

where $D$ is the scatter plot box-counting dimension. $D$ is then defined by the slope in this linear regime, that is,

$$N(\epsilon) \propto \epsilon^{-D}$$

We estimate the error in $D$ through the least squares fit for the slope.

Further details and discussion on the method can be found in Ref. Domínguez et al. (2014).

It is clear that, in order to calculate $D$, a certain time frame of the dataset must be chosen. Given the time windows chosen for $Dst$

The method as stated above was applied to the $Dst$ time series where, given the width of the data windows used (the criterion is discussed in Sec. 3) and the time resolution of the data (one point per hour), it only made sense to build the scatter plot with consecutive data points.

However, when resolution is larger, as is the case with simulation and solar wind data, it is possible to consider different time delays. Thus, the scatter plot can be built by plotting the $i$-th data in the set, versus $de (i+j)$-th data, with $j \geq 1$ in general, and then the fractal dimension calculated depends on $j$, $D_j$. This was the approach in Refs. Domínguez et al. (2017) and Muñoz et al. (2016), and presented here in Secs. 4 and 5.
3 Dst time series: Storm and quiet states

We first apply this technique to quiet and active periods with magnetic storms in order to investigate the relationship between the intensity of the Some studies (Balasis et al., 2009; Papa and Sosman, 2008) have suggested that there is a relationship between the intensity and the complexity of the Dst index and its fractal dimension, a relationship which has also been suggested by other studies of the complexity time series. Here, we will first apply the technique discussed in Sec. 2 to investigate whether there is a connection between the level of geomagnetic activity and the fractal dimension of the Dst series. (Balasis et al., 2009; Papa and Sosman, 2008) index.

Following Ref. Domínguez et al. (2014), we identify “storm states” and “quiet states” by locating peaks in the Dst series where peaks, such that $D_{st} < -220$ nT, and then a “storm state” is defined by a window starting one week before the minimum value of the peak, and ending one week after it. This is done considering the typical time scale of a geomagnetic storm (Tsurutani and Gonzalez, 1994; Gonzalez et al., 1994). Then, the “quiet state” corresponds to the period of time between two “Quiet states” simply correspond to the time window between consecutive “storm states”. Figure 2 illustrates this by showing the four peaks detected in 1989 and the corresponding windows, and where four peaks are found.

![Figure 2. Storm and quiet states in the Dst time series for 1989. Identifying the storm and quiet states as explained in Sec. 3. The solid line shows also indicating the average value of the Dst index (horizontal line), and the threshold value used to identify a geomagnetic storm. Red dots show the minimum Dst value used to identify a “storm state” storms (dashed line). Red and (black) arrows show windows corresponding to indicate storm and (quiet) states, respectively. (Taken from Domínguez et al. (2014).)](

For future identification, we label each state in a year with consecutive integer numbers. In the following, states within a year are label by integer numbers starting from 1. For instance, in Fig. 2, the year starts with a quiet state, then that will be state “1”, the following state will be a storm, and it will be state “2”. Thus, all future quiet states within the year will be labeled with consecutive odd numbers, whereas storm states will be labeled with consecutive even numbers.

The box-counting dimension A fractal dimension is then calculated for each storm and quiet state, calculated as described in Sec. 2, is each quiet state in the same year. Results for 1989 are shown in Fig. 3. Red circles indicate storm states. Error bars in $D$ are given by the error of its least squares linear fit.

Similar plots for 5 years of high geomagnetic activity were obtained (Domínguez et al., 2014). In general, it is found that storm states storm states are found to have smaller fractal dimension than the surrounding quiet states quiet states immediately before and after them, although there does not seem to be a clear correlation on the value of Dst itself, and the fractal dimension,
as shown in Fig. 22 for all states, for all years studied in Domínguez et al. (2014). No obvious correlation is found if individual years are considered either (Domínguez et al., 2014). Thus, our statement on the decrease of the fractal dimension is an argument on its variation, rather than on its actual value.

Mean value of $D_{st}$ for each state as function of the box counting dimension $D$ with respective error bars (calculated as in Fig. 3), for five years of high geomagnetic activity: 1960, 1989, 2000, 2001, and 2003.

4 $D_{st}$ time series: Variable width windows around a storm

We have also studied variable width windows around a storm and moving windows across storms, and results have been consistent with the findings discussed.

If the qualitative connection between fractal dimension and existence of a storm observed in Sec. 3 is robust, then widening the window. In effect, as a window is widened around a storm should increase its fractal dimension, as more “quiet” data are taken into account.

To this end, we take windows starting/ending $n$ weeks before/after the peak, with $n = 1, \ldots, 6$. We illustrate this considered, and thus the fractal dimension of the data inside the window should increase. This is actually the case, as shown for instance in Fig. 22, where the windows considered around the 13 March 1989 storm are shown.

Variable size windows around the 13 March 1989 storm (peak at abscissa 1729). The plot shows the $D_{st}$ index as a function of time, measured in hours since the beginning of the year.

Figure 4(a) shows the 4, where results for four particular storms: 1 April 1960, 13 March 1989, 6 April 2000, and 30 March 2001, with minimum intensities of $-327$ nT, $-589$ nT, $-288$ nT, and $-287$ nT, respectively. These storms have been chosen because they are isolated, so that windows can be enlarged (up to four weeks on each side) without including new “storm states” enough to allow enlarging of the window around them without overlapping with neighboring storms (Domínguez et al., 2014).

Figure 5 shows the results for the fractal dimension. Regarding the moving windows analysis, results are illustrated for the 13 March 1989 storm in Fig. 5, comparing it with the $D_{st}$ index.
Figure 4. (a) Box-counting dimension $D$ for a storm state with respective error bars, as a function of the width of the window around $D_{st}$. (b) Mean value of $D_{st}$ for each variable width window around the same Linetns correspond to storms in (a) on 1 April 1960, as function of the box-counting dimension with respective error bars. 13 March 1989, 6 April 2000, and 30 March 2001. (Taken from Domínguez et al. (2014).) Consistent with the results in Sec. 3, the box-counting dimension increases as we zoom out from the storm, which means that the relevance of the storm itself within the window decreases. This is confirmed by plotting the mean value of $D_{st}$ in a window as a function of $D$, for the same storms. This is shown in Fig. 7(b). We expect that increasing the window width should increase not only the value of $D$ as noted above, but also the average value of $D_{st}$ for the same reason, and thus $D$ and $\langle D_{st} \rangle$ should be positively correlated. This is confirmed in Fig. 7(b). The breaks in the linear behavior for some curves can be explained by the existence of nearby peaks close to the storm studied, as explained in detail in Domínguez et al. (2014).

4 $D_{st}$ time series: Moving windows across a storm

We now calculate the fractal dimension for fixed-width windows (two weeks), initially placed well before the storm peak, and move it in steps of one week crossing the peak. This will give us a better intuition on the evolution of the fractal dimension in time, in particular during a storm. The initial position of the window is the first day of the year, and it is moved until it reaches the third week after the peak (see Fig. 22).

Figure 5. Box-counting dimension $D$ (blue, with error bars) and $D_{st}$ index (red) for the 13 March 1989 geomagnetic storm. Vertical lines show windows of data where $D$ decreases before the storm in the windows within the vertical lines. (Taken from Domínguez et al. (2014).)

For all cases studied (Domínguez et al., 2014) As shown in Domínguez et al. (2014), the box-counting dimension of the $D_{st}$ index decreases as the storm approaches. However, it is very interesting to note that we have a noticeable change in the fractal dimension, even for all cases studied. Moreover, this decrease occurs before the window contains any point of increased geomagnetic storm. This is illustrated, as marked by the vertical lines in Fig. 5, where two vertical lines indicate the window...
of $Dst$ immediately before the storm. The storm is not included in the window, however the fractal dimension has already started to decrease. Whether this is relevant for forecasting geomagnetic storm needs further study, as it may simply be due to an increase of the intermittency in the time series, unrelated to the upcoming dissipative event.

In Domínguez et al. (2014), systematic calculations of cross correlation between $Dst$ and $D$ were performed for all storms analyzed, and for the same five complete years studied in that paper (using year-long data for 1960, 1989, 2000, 2001, and 2003), which have already been analyzed, but only near geomagnetic storms. Results suggest that the box counting dimension consistently decreases when the storm approaches, thus suggesting that the decrease of the box-counting dimension of the $Dst$ series, or similar measures of complexity, could be of relevance when forecasting geomagnetic storms. Dimension is a robust feature.

We also studied the possible correlation between the fractal dimension and measures of solar activity, to investigate whether this simple measure of complexity yields any information about the connection between solar and geomagnetic activities. In particular, we considered the solar flare index (Atıcı and Özgüc, 1998; Özgüc et al., 2003) and the coronal index (Rybansky et al., 2001; Na, which are measures of energy released from the Sun.

Results are shown in the left panel of Fig. 22 for the solar flare index, and in the right panel of the same figure for the coronal index, using a moving windows approach over the 13 March 1989 storm, the same storm we have described in the previous sections. Similar analyses were performed for events in 2000 and 2001, as shown in Domínguez et al. (2014). It is found that even for solar flare events of different intensities, periods of large solar flare index are accompanied by a decrease in the fractal dimension $D$ of the $Dst$ time series. In the case of the coronal index, results suggest that one or two weeks before the minimum value of $D$, which corresponds to the storm, there is a maximum in the coronal index. However, this is only clearly seen regarding positions of maximum/minimum values. A more detailed correlations analysis, using daily coronal index data does not show any particular signature.

-Box counting dimension $D$ (with error bars) corresponding to the $Dst$ index, along with the total solar flare (sum of northern and southern hemispheres indexes) (left panel) and coronal (right panel) indexes for the storms: 13 March 1989, with moving windows.

We observe that two different estimations of solar activity are correlated to some extent with $D$, thus suggesting a link between the solar activity and the fractal features of the Earth’s magnetosphere. Certainly, one should probably not expect to find a single index to reveal this, as geomagnetic dynamics may be mostly but not exclusively determined by solar behavior, and several other correlated pairs have been proposed (Yurchyshyn et al., 2004), but it is interesting to notice the overall consistency of the results, at least when a correlation can be observed.

## 4 MHD Shell Model

Given the intrinsic difficulties in using direct numerical simulations to describe turbulent flows, specially for large Reynolds numbers, shell models have been used for years in order to reproduce the nonlinear dynamics of fluid systems in large dynamical ranges, but with less degrees of freedom (Obukhov, 1971; Gledzer, 1973; Yamada and Ohkitani, 1988). An MHD shell
model (Boffetta et al., 1999), in particular, is a dynamical system which aims to reproduce the main features of MHD turbulence. The model corresponds to a simplified version of the Navier-Stokes or MHD equations for turbulence, that conserves some of its invariants in the limit of no dissipation.

In this work, we use the MHD GOY shell model, which describes the dynamics of the energy cascade in MHD turbulence (Lepreti et al., 2004). The model is built up by dividing the wave-vector space \((k\text{-space})\) in \(N\) discrete shells of radius \(k_n = k_0 2^n\) \((n = 0, 1, \ldots, N)\). Then, two complex dynamical variables \(u_n(t)\) and \(b_n(t)\) representing velocity and magnetic field increments on an eddy scale \(l \sim k_n^{-1}\), are assigned to each shell.

The model consists of the following set of ordinary differential equations:

\[
\begin{align*}
\frac{du_n}{dt} &= -\nu k_n^2 u_n + ik_n \left( u_{n+1} u_{n+2} - b_{n+1} b_{n+2} \right) - ik_n \left\{ \frac{1}{4} \left( u_{n-1} u_{n+1} - b_{n-1} b_{n+1} \right) + \frac{1}{8} \left( u_{n-2} b_{n-1} - b_{n-2} u_{n-1} \right) \right\}^* + f_n, \\
\frac{db_n}{dt} &= -\eta k_n^2 b_n + ik_n \frac{1}{6} \left( u_{n+1} b_{n+2} - b_{n+1} u_{n+2} \right) - ik_n \frac{1}{6} \left\{ \left( u_{n-1} b_{n+1} - b_{n-1} u_{n+1} \right) + \left( u_{n-2} b_{n-1} - b_{n-2} u_{n-1} \right) \right\}^* + g_n,
\end{align*}
\]

(2)

where \(\nu\) and \(\eta\) are, respectively, the kinematic viscosity and the resistivity; \(f_n\) and \(g_n\) are external forcing terms acting, respectively, on the velocity and magnetic fluctuations. The nonlinear terms have been obtained by imposing quadratic nonlinear coupling between neighbouring shells and the conservation of three MHD ideal invariants (Gloaguen et al., 1985; Lepreti et al., 2004).

The forcing terms are calculated according to the Langevin equation

\[
\frac{df_n}{dt} = -\tilde{f}_n \tau_0 + \tilde{\mu},
\]

(4)

where \(\tilde{f}_n = f_n\) or \(g_n\), \(\tau_0\) is a characteristic time of the largest shell and \(\tilde{\mu}\) is a Gaussian white noise of width \(\sigma\).

The magnetic energy dissipation rate is defined as

\[
\epsilon_b(t) = \eta \sum_{n=1}^{N} k_n^2 |b_n^2|.
\]

(5)

In our simulation, we set \(\sigma = 0.01\), \(\tau_0 = 0.25\), take \(N = 19\) shells, and force the system on the largest shell \((f_1, g_1 \neq 0)\). Similar parameters have been considered in previous studies using this model for modelling of solar flares statistics (Boffetta et al., 1999; Lepreti et al., 2004; Nigro et al., 2004).

We numerically integrate the shell model Eqs. (2)–(3) for various values of \(\nu\) and \(\eta\), and then we calculate the magnetic energy dissipation rate \(\epsilon_b(t)\) (Eq. (5)).

Figure 6 shows a typical time behavior for \(\epsilon_b(t)\).

Previous works have compared the statistics of bursts in turbulent systems with the statistics of dissipative events in the shell model (Boffetta et al., 1999; Lepreti et al., 2004; Carbone et al., 2002). There, peaks in the \(\epsilon_b(t)\) time series have been associated to dissipative events in the magnetized plasma. Following the ideas in Ref. Domínguez et al. (2014) .
we focus only on the largest peaks in the $\epsilon_b(t)$ time series, specifically, on dissipative events where the maximum value is larger than $\langle \epsilon_b \rangle + n \tilde{\sigma}$ where $\langle \epsilon_b \rangle$ is the average value of $\epsilon_b$ over all simulation time, $\tilde{\sigma}$ is the standard deviation of the $\epsilon_b$ time series in that window, and $n$ is a certain integer. In this paper we discuss only results for $n = 10$, but in Ref. Domínguez et al. (2017) Domínguez et al 5 was also considered, in order to assess the robustness of the results. Our aim is to study the dependence of the conclusions on $\nu$ and $\eta$ in Eqs. (2) and (3).

5 Active and quiet states in the MHD shell model

We now apply the same techniques used to study the $Dst$ index, as described in Secs. 3–5, to the $\epsilon_b(t)$ time series.

First, we first notice that, in general, setting parameters $\nu$ and $\eta$ with arbitrary values yields $\epsilon_b(t)$ series which do not have the necessary intermittency level to resemble the $Dst$ time series. Compare, for instance, the different panels in Fig. 6 in Domínguez et al. (2017), which shows that $Pm = 0.2$ leads to a very noisy output, unlike simulations with $Pm = 1.0$ or 2.0, where individual, large peaks can be easily identified from the background. In fact, previous studies have shown that the statistics of bursts follows a power law for $Pm = 1$ (Boffetta et al., 1999; Lepreti et al., 2004; Carbone et al., 2002), and for this reason we start by taking $Pm = \nu/\eta = 1$.

Now, we need to define “active states” and “quiet states”. However, unlike For the $Dst$ case (Domínguez et al., 2014) there is no clear criterion to establish the time scale of a typical dissipative event for our simulation data, and thus we proceed there are natural time scales for which were used to define the occurrence and the duration of a geomagnetic storm. However this is not available for the output of the shell model, and our approach was to inspect the data. To this end, and in order to explore a wide range of parameters, we fix $Pm = \nu/\eta = 1$, and take values $\nu = \eta = 10^{-\mu}$ with $\mu = 1, 2, 3, \ldots, 12$. We then as fixed, and then a wide range of values of $\nu$, in the interval $10^{-12} \leq \nu \leq 10^{-1}$. For each parameter set, we solve the shell model equations with using a time step of $dt = 10^{-4}$ and for $7 \times 10^8$ iterations. This series of simulations suggest We conclude that $n = 10$ is enough to identify the largest peaks, filtering out most of the other events filter most events, except for the largest ones.
Regarding the width of an active states, Fig. 6 is, among the various simulations we performed, the only case where two clear dissipative events were both close and distinguishable from each other. Thus, we take this as a reference and we define an active state width such that both peaks in Fig. 6 can be regarded as two separate events. Since the separation between both peaks is 96,000 time steps, we define an “active state” by identifying a peak, and then considering a window starting 48 in the shell model output as a window of 96,000 time steps before, and ending 48,000 time steps after it. With this definition, an active state centered around a peak in the magnetic energy dissipation. Therefore, Fig. 6 shows two adjacent active states, each one associated with one of the peaks.

With these definitions of active and quiet states, we analyze the simulation results for \( \nu = \eta = 10^{-\mu} \) with \( \mu = 3 \) and calculate the scatter box-counting dimension for each state for various values of the sampling \( j \). Figure 7 shows the results for \( \nu = \eta = 10^{-3}, n = 10 \). Three quiet states and two active states are identified. They are identified by integer numbers following the same strategy described in Sec. 3.1. In Fig. 7, active states correspond to labels “2” and “4” corresponding to the active states.

Performing the procedure described in section 2, we calculate the scatter box-counting dimension for different values of \( j \) for each active and quiet state. Results are shown in Figure 7. Errors bars in \( D \) are given by the error of calculated from the least squares linear fit.

![Figure 7](image)

**Figure 7.** Box counting fractal dimension for \( \epsilon_b(t) \) during quiet and active states for \( n = 10 \), with \( \mu = 3 \). Active states correspond to states labeled “2” and “4”. (Taken from Domínguez et al. (2017).)

We note that in general, an active state has a smaller fractal dimension than the surrounding quiet states. This is observed for all values of \( j \) considered, although quantitative differences occur. For instance, in Fig. 7 we note that when \( j \) decreases, for all values of \( j \), we notice that active states have smaller fractal dimensions than neighbouring quiet ones, although the difference between quiet and active states is less clear.

Figure 7 also shows that the fractal dimension depends on the distance between consecutive data, represented by the value of the sampling parameter \( j \), which may be seen as an indication of an underlying multifractal structure that suggests multifractal features of the data (Kadanoff et al., 1995; Pisarenko et al., 1993).
In order to further investigate the dependence on $j$, we plot We follow this idea by plotting the fractal dimension for each quiet and active state in the simulations of each state as a function of the distance between data, $j$. Results are shown in (see Fig. 8).

Figure 8. Box counting fractal dimension for quiet and active states, as a function of $j$, with $\mu = 3$. Numbers for each curve rotate the states. We added an “(s)” in the legends, in order to highlight the active states. (Taken from Domínguez et al. (2017).)

As mentioned above, the scatter box fractal dimension when all data are taken Notice that for the smallest value of $j$ ($j = 1$), the scatter plot is a straight line, yielding thus its fractal dimension is $D = 1$, consistent with Fig. 8. On the other hand, as $j$ increases, a smaller subset of simulation data is taken, and eventually, when $j$ and is larger than the number of data, only one datum is taken, leading to $D = 0$ points, $D = 0$, since such a sampling leaves only one point in the curve. Both limits are found satisfied for all curves. For in Fig. 8. The nontrivial dependence of $D$ for intermediate values of $j$, a nontrivial dependence of the fractal dimension is observed, which also reflects the reflects, as mentioned above, a multifractal nature of the series as $D$ varies as we change the time scale given by $j_{\varepsilon(t)}$ times series, since the fractal dimension depends on the sampling time scale.

Figure 8 shows that active states have lower fractal dimensions than quiet states, consistently with Fig. 7. Moreover, active states always have fractal dimensions less than 1, whereas it is always larger than 1. As observed in Fig. 7, in Fig. 8 we also find lower fractal dimensions for active state than for quiet states. It also shows that all quiet (or equivalently all active) states are not characterized by a single fractal dimension, consistent with our previous findings. Finally, Fig. 8 shows that the fractal dimension decreases during dissipative events for a certain range. However, this does not hold for arbitrarily large values of $j$. If (see Figs. 6 and 7 in Domínguez et al. (2017)). We only find it for $j > 1$, but not too large, the fractal dimension during active states is always smaller than during quiet times, which suggests that, for suggesting that it is within this range of moderate values of $j$, where the box counting fractal dimension has statistical information on the activity of dissipative events in the time series. This, in turn, suggests that the findings in Sec. 3 and Domínguez et al. (2014) are not trivial.

In Ref. Domínguez et al. (2017) a more detailed analysis is carried out on the shell model results, exploring other simulation parameters ($\nu$, $\eta$, magnetic Prandtl number), other criterion for defining active states ($n = 5$), and a systematic study of the correlations between the fractal dimension and the occurrence of dissipative events by means of the Student’s $t$-test. Results suggest that the intermittency level of the output time series is relevant, which has led us to perform the analyses for the shell model within a certain range of values of the Prandtl number, as well as of the viscosity and resistivity.

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5 Magnetic clouds

As a way to illustrate how the ideas described so far could be used to characterize structures in space plasmas, we apply the method to study the time series for the magnetic field during magnetic clouds (Burlaga et al., 1981), as found in ACE data (ACE Science Center). Magnetic clouds are transient structures ejected from the Sun, characterized by a large and smooth rotation of the magnetic field. Typically, a magnetic cloud event can be identified from single spacecraft measurements by studying the evolution of the observed fields. During a given event, various stages can be identified: first, observation of solar wind prior to the cloud’s arrival, then a sheath of compressed solar wind plasma immediately preceding a flux rope, where the magnetic field varies smoothly, and finally the background solar wind again. Note that slower-moving clouds traveling at speeds comparable to that of the ambient solar wind will not display prominent sheath regions.

Two events were selected: an event occurring on 12 July 2012 (MC1) and another on 11 July 2014 (MC2). Resolution for the magnetic field time series for this event is 16 seconds, covering a time span of 8 days for MC1 and 6 days for MC2, of which about 2 days correspond to the cloud event itself. It is found that the calculated fractal dimension evolves in a distinctive way as the various stages of the event as it passes by the spacecraft (namely surrounding solar wind, sheath, and flux rope). Given the high resolution of the data, it is possible to calculate the box-counting dimension for several delays, given by $j$, as was shown in Figs. 7 and 8 for the shell model analysis. In Fig. 9 the fractal dimension is calculated for each magnetic cloud stage, and various values of the sampling $j$ are considered.

![Figure 9](image_url)

**Figure 9.** Box counting fractal dimension for two magnetic cloud events during the four stages of the time series: first the solar wind, then the sheath, then the flux rope, and finally the solar wind again. Several values for data sampling $j$ are used.

It can be noted that the fractal dimension, as calculated here, is indeed able to characterize magnetic cloud structures. The sheath state has a large dispersion of fractal dimension values as $j$ is varied, consistent with its more turbulent regime; on the other hand, the quieter and more organized flux rope state exhibits a very low variation with $j$, basically a single fractal
dimension at all time scales explored. As for the surrounding solar wind, it shows dispersion of $D_j$ which is between the dispersion of values in the sheath and the flux rope (Muñoz et al., 2016).

6 Conclusions

In this manuscript, we have reviewed recent results obtained by us, regarding the evolution of complexity in magnetized plasmas, as described by geomagnetic data, simulation results for MHD turbulence, and spacecraft data in the solar wind.

This has been done by calculating a box-counting fractal dimension for time series of magnetic field data for the $Dst$ geomagnetic index (Domínguez et al., 2014), the GOY shell model (Domínguez et al., 2017), and ACE data for two magnetic cloud events (Muñoz et al., 2016).

Some robust behaviors are identified. In general, it is found that the fractal dimension $D$ decreases during dissipative events.

In the case of the $Dst$ time series this was verified for three different types of time windows: fixed width and stationary (Sec. 3), variable width (Sec. ??), and moving windows (Sec. 5). And it was also found across several time scales, namely individual storms, full years, and the complete 23rd solar cycle, as detailed in Ref. Domínguez et al. (2014).

A similar behavior is found for the MHD shell model (see Secs. 4 and 5). Thanks to the larger resolution of the simulation data as compared with the $Dst$ data, several values of the time delay for data sampling could be made, showing that the results found in Ref. Domínguez et al. (2014) are nontrivial, in the sense that not all samplings yield similar results. Only intermediate, not too large, values of the time delay (as represented by the value of $j$ in Sec. 5) are able to clearly distinguish between active and quiet states. But, within the useful range of values for $j$, the fractal dimension of the active states is consistently smaller than the dimension of quiet states, and is always lower than 1, whereas the active states always have a dimension larger than 1. The dependence on $j$ of the fractal dimension is interesting in itself, as it suggests that data have a multifractal structure, which is consistent with suggestions and finding by other authors for space plasmas (Chapman et al., 2008, 1998; Valdivia et al., 2005).

Also, a more systematic test for the correlation between burst events in the shell model and the decrease in fractal dimension was performed, by means of the Student’s $t$-test, as well as a more detailed exploration of the parameter space for the simulation. These results can be found in Ref. Domínguez et al. (2017) .

As an application of these ideas, we take two magnetic cloud events in the solar wind, and use the techniques described here to study the corresponding magnetic field time series. Our results, although preliminary, suggest that this method can characterize the various stages of the magnetic cloud structure.

Given the rich and complex dynamics governing the evolution of magnetized plasmas, we would not expect that a single index would be able to capture all their relevant information. In fact, multifractal analysis should be made in order to represent the dynamics of the systems studied more accurately, and such an analysis is currently being prepared for future publication. However, the findings summarized here suggest that some relevant correlations can be observed, and that the dimension used here, although simple, may give some insight on the evolution of complexity of plasmas in the Sun-Earth system and MHD turbulent states.
Competing interests. The authors declare that they have no conflict of interest.

Acknowledgements. This project has been financially supported by FONDECYT under contracts Nos. 1110135, 1110729, and 1130273 (J.A.V.); Nos. 1080658, 1121144, and 1161711 (V.M.); and No. 3160305 (M.D.). M.D. also thanks a doctoral fellowship from CONICYT, and a Becas-Chile doctoral stay, contract No. 7513047. We are also thankful for financial support by CEDENNA (J.A.V.), and the US AFOSR Grant FA9550-16-1-0384 (J.A.V. and V.M.).
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